

Sensitivity analysis for optimization under constraints and with uncertainties

Sensitivity analysis on (excursion) sets

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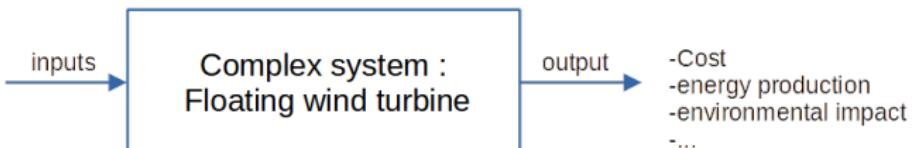
École Centrale de Lyon & IFP Énergie nouvelles

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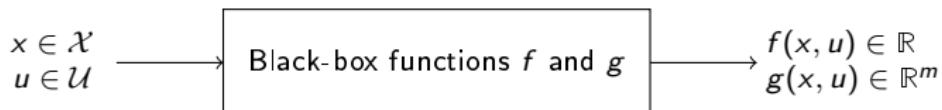


System & black-box model

Input parameters :
-component size
-material
-swell height
-...



System & black-box model



- The x are the deterministic inputs
- The u are uncertain inputs : $u = U(\omega)$ with U a random vector of density ρ_U
- f is the objective function to minimize
- g is the constraint function defining the constraint to respect : $g \leq 0$

Optimization problem

Robust optimization problem

$$\begin{aligned} x^* = \arg \min_x & \mathbb{E}[f(x, U)] \\ \text{s.t. } & \mathbb{P}[g(x, U) \leq 0] \geq P_{\text{target}} \end{aligned} \tag{1}$$

Deterministic strategy

$$\begin{aligned} x^* = \arg \min_x & F(x) \\ \text{s.t. } & G(x) \leq 0 \end{aligned} \tag{2}$$

- Goal Oriented Sensitivity Analysis to reduce the input dimension : Spagnol 2020
- Huge evaluation cost of F and G
- What about the U ?

How to quantify the impact of the uncertain inputs U on the optimization ?

Toy function

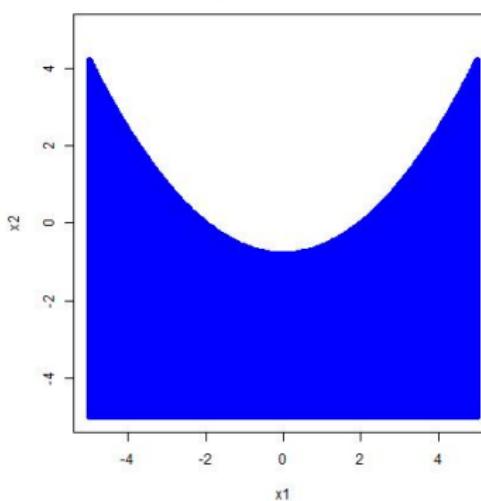
Toy function from [El-Amri et al. 2021]

$$\forall (x, u) \in [-5, 5]^4 \quad g(x, u) = -x_1^2 + 5x_2 - u_1 + u_2^2 - 1$$

$$\Gamma_u = \{x \in [-5, 5]^2, g(x, u) \leq 0\}$$

u_2 fixé

$u_1 = -5 \quad u_2 = 0$



Toy function

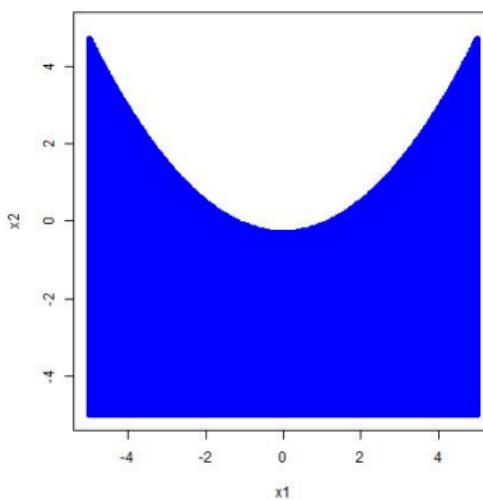
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u_2 fixé

u1 = -2.5 u2 = 0



Toy function

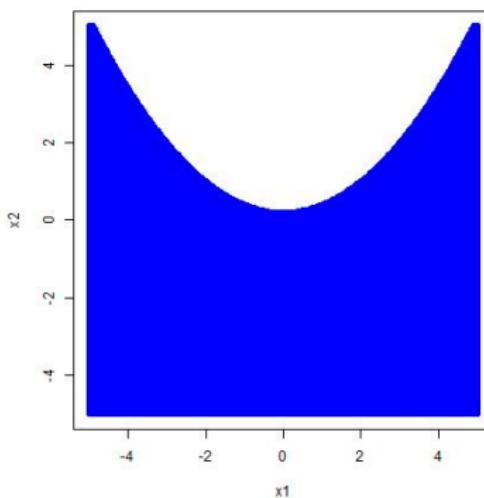
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$$\Gamma_u = \{x \in [-5, 5]^2, g(x, u) \leq 0\}$$

u_2 fixé

$u_1 = 0 \quad u_2 = 0$



Toy function

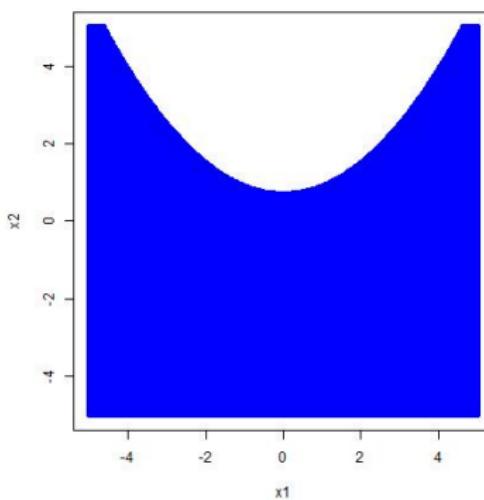
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$$\Gamma_u = \{x \in [-5, 5]^2, g(x, u) \leq 0\}$$

u_2 fixé

$u_1 = 2.5 \quad u_2 = 0$



Toy function

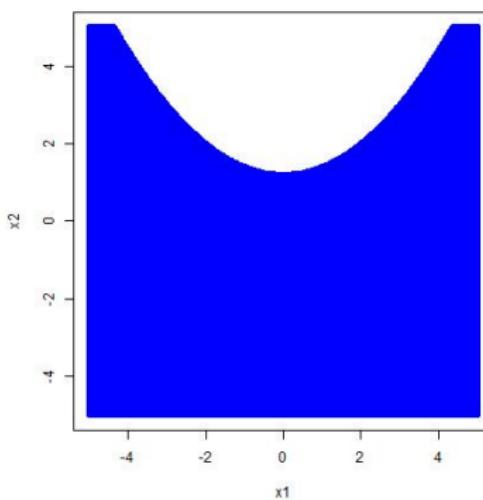
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$$\Gamma_u = \{x \in [-5, 5]^2, g(x, u) \leq 0\}$$

u_2 fixé

$u_1 = 5 \quad u_2 = 0$



Toy function

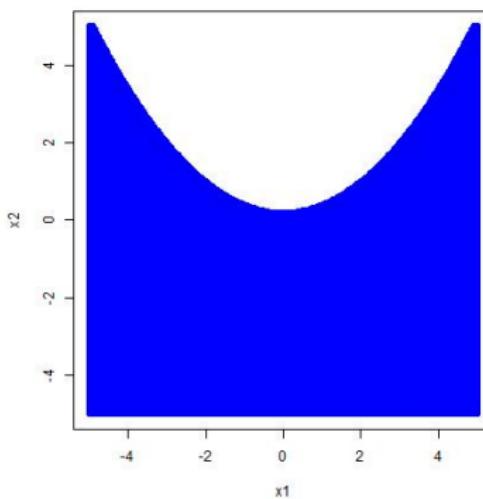
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u_1 fixé

$u_1 = 0 \quad u_2 = 0$



Toy function

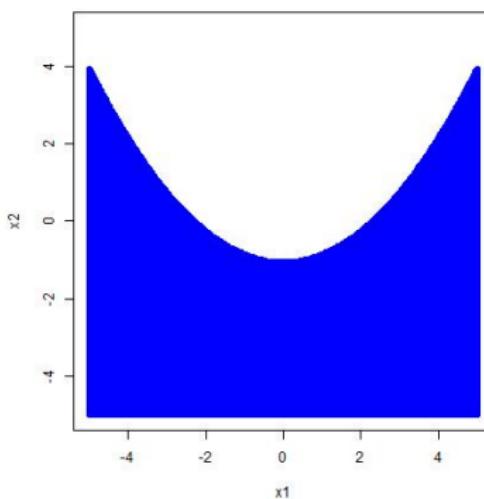
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$$\Gamma_u = \{x \in [-5, 5]^2, g(x, u) \leq 0\}$$

u_1 fixé

$u1 = 0 \quad u2 = 2.5$



Toy function

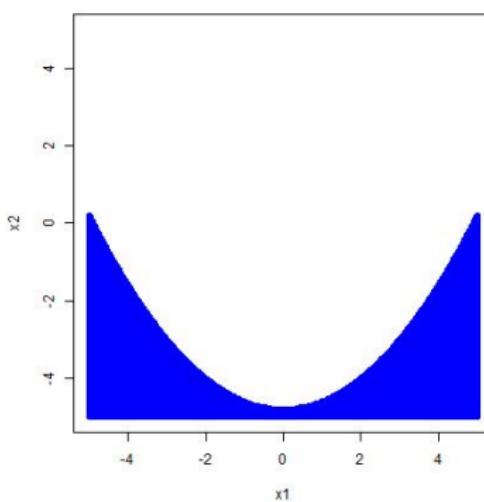
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$$\Gamma_u = \{x \in [-5, 5]^2, g(x, u) \leq 0\}$$

u_1 fixé

$u_1 = 0 \quad u_2 = 5$



Subproblem : Excursion sets

Excursion sets

New output :

$$\Gamma_U = \{x \in \mathcal{X}, g(x, U) \leq 0\}, \quad (3)$$

which is called a random excursion set.

Influence of the uncertain inputs U on Γ_U ? \Rightarrow SA on excursion sets.

Yet most SA techniques have scalar or vectorial outputs.

How then can we do sensitivity analysis on (excursion) sets ?

- SA on sets using random set theory notably Vorob'ev expectation and deviation
- SA on sets using universal indices from Gamboa et al. 2021
- **SA on sets using RKHS theory**

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- Distribution embedding into a RKHS
- From a distance between distributions to sensitivity indices

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- A kernel between sets
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Sobol indices on random sets ?

- How to quantify the influence of the input U_i on the output $\Gamma_U = \{x \in \mathcal{X}, g(x, U) \leq 0\}$ which is a random set.
- "Sobol indices" on random sets :

$$S_i = \frac{\text{Var } \mathbb{E}[\Gamma|U_i]}{\text{Var } \Gamma}$$

→ Expectation and Variance of a random set ?

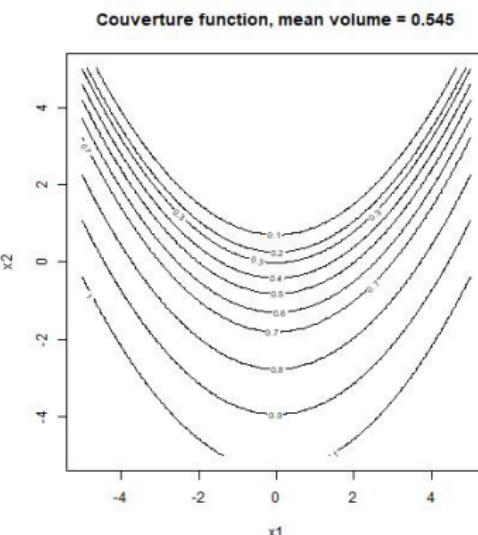
Vorob'ev expectation and deviation (Molchanov 2005)

Vorob'ev expectation

$$\mathbb{E}^V[\Gamma] = \{x \in \mathcal{X}, \mathbb{P}(x \in \Gamma) \geq \alpha^*\} = Q_{\alpha^*} \text{ with } \mathbb{E}[\mu(\Gamma)] = \mu(Q_{\alpha^*}) \quad (4)$$

Vorob'ev deviation

$$\text{Var}^V(\Gamma) = \mathbb{E}[\mu(\Gamma \Delta \mathbb{E}^V[\Gamma])], \quad (5)$$



Vorob'ev expectation and deviation (Molchanov 2005)

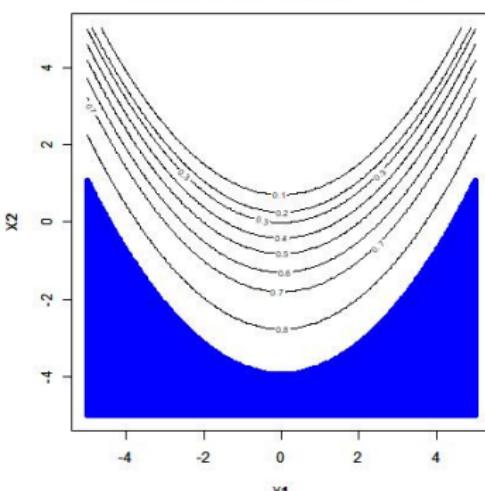
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Couverture function, mean volume = 0.55
 Quantile of order 0.9, volume = 0.27



Vorob'ev expectation and deviation (Molchanov 2005)

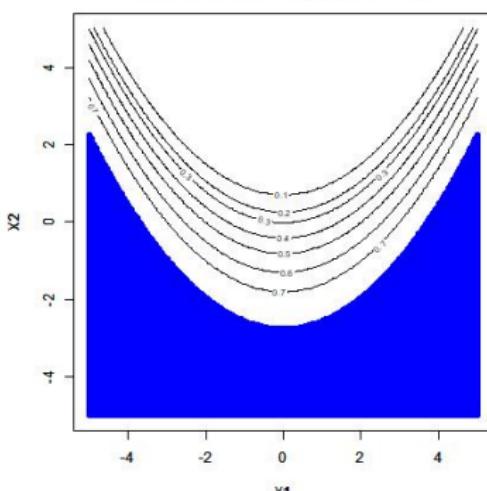
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$$\text{Var}^V(\Gamma) = \mathbb{E}[\mu(\Gamma \Delta \mathbb{E}^V[\Gamma])], \quad (5)$$

Couverture function, mean volume = 0.55
 Quantile of order 0.8, volume = 0.39



Vorob'ev expectation and deviation (Molchanov 2005)

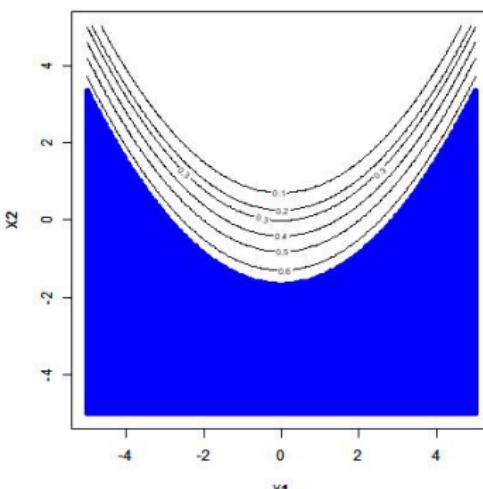
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$$\text{Var}^V(\Gamma) = \mathbb{E}[\mu(\Gamma \Delta \mathbb{E}^V[\Gamma])], \quad (5)$$

Couverture function, mean volume = 0.55
 Quantile of order 0.7, volume = 0.5



Vorob'ev expectation and deviation (Molchanov 2005)

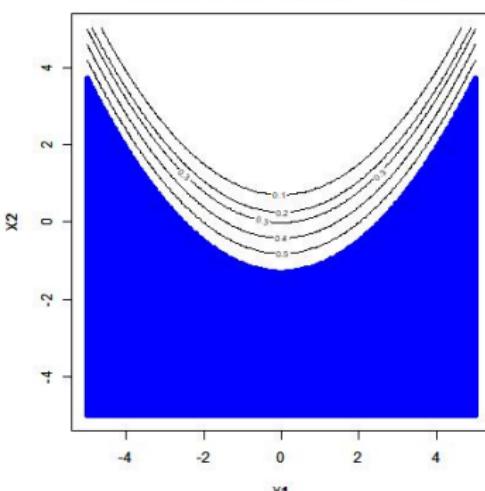
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Vorob'ev deviation

$$\text{Var}^V(\Gamma) = \mathbb{E}[\mu(\Gamma \Delta \mathbb{E}^V[\Gamma])], \quad (5)$$

Couverture function, mean volume = 0.55
 Quantile of order 0.6, volume = 0.54



Vorob'ev expectation and deviation (Molchanov 2005)

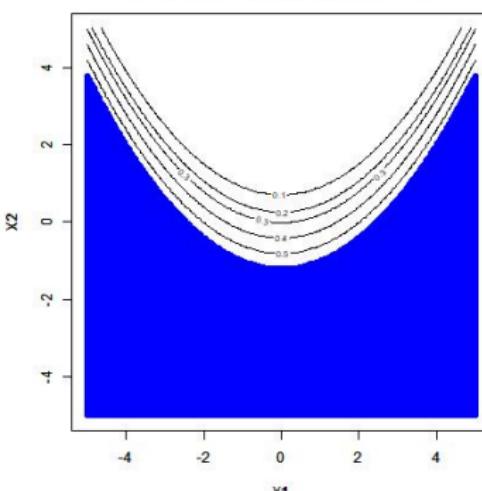
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Vorob'ev deviation

$$\text{Var}^V(\Gamma) = \mathbb{E}[\mu(\Gamma \Delta \mathbb{E}^V[\Gamma])], \quad (5)$$

Couverture function, mean volume = 0.55
 Quantile of order 0.57, volume = 0.55



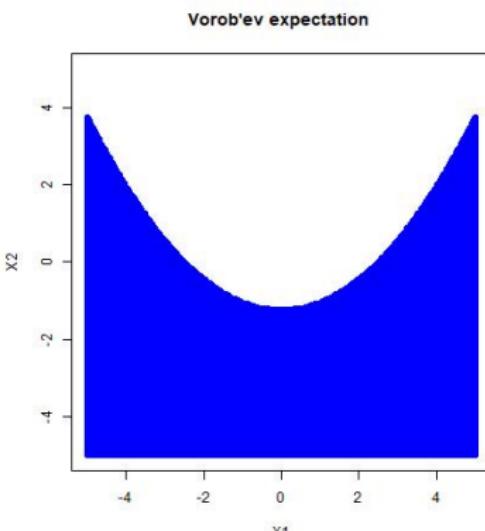
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Vorob'ev deviation

$$\text{Var}^V(\Gamma) = \mathbb{E}[\mu(\Gamma \Delta \mathbb{E}^V[\Gamma])], \quad (5)$$



Definition of the proposed indices & estimation

First order Vorob'ev index :

$$\begin{aligned} S_i^V &= \frac{\text{Var}^V(\mathbb{E}^V(\Gamma|U_i))}{\text{Var}^V(\Gamma)} \\ &= \frac{\mathbb{E}[\mu(\mathbb{E}^V(\Gamma|U_i)\Delta\mathbb{E}^V[\mathbb{E}^V(\Gamma|U_i)])]}{\mathbb{E}[\mu(\Gamma\Delta\mathbb{E}^V(\Gamma))]} . \end{aligned}$$

But :

$$\mathbb{E}^V[\mathbb{E}^V(\Gamma|U_i)] \neq \mathbb{E}^V(\Gamma)$$

which has two mains issues :

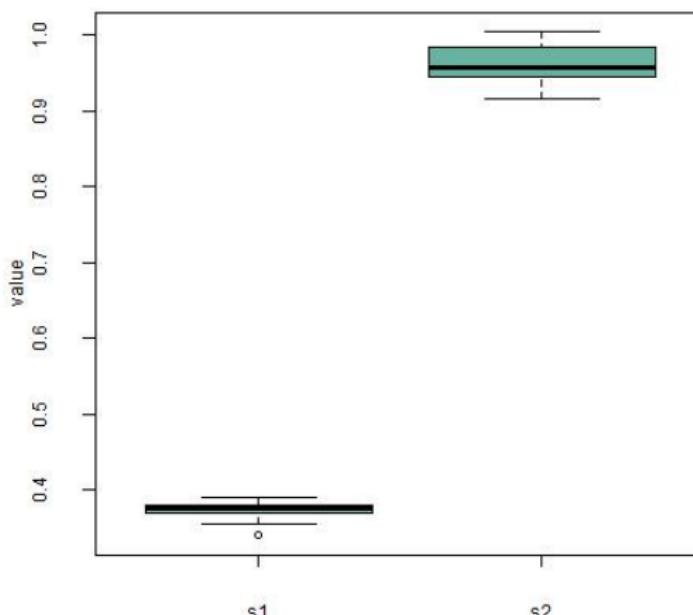
- Very costly Monte Carlo estimation
- No variance decomposition

Results

Toy function

$$\forall (x, u) \in [-5, 5]^4 \quad g(x, u) = -x_1^2 + 5x_2 - u_1 + u_2^2 - 1$$

Vorob'ev indices for N_x=20, N_u=500



Adaptation of the universal index on random sets

$$S_{2, \text{ Univ}}^i(T_a, \mathbb{Q}) := \frac{\int_{\mathcal{A}} \text{Var}(\mathbb{E}[T_a(\Gamma_U) | U_i]) d\mathbb{Q}(a)}{\int_{\mathcal{A}} \text{Var}(T_a(\Gamma_U)) d\mathbb{Q}(a)}$$

We use $T_a(\Gamma)$ defined by :

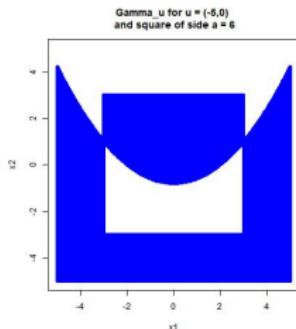
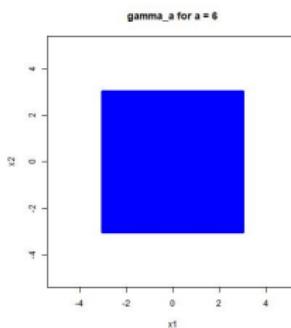
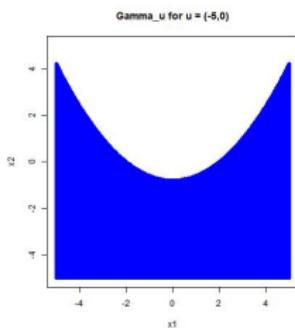
$$T_a(\Gamma) = \mu(\gamma_a \Delta \Gamma), \quad (6)$$

with :

- the symmetric difference Δ defined by $A \Delta B = A \cup B - A \cap B$.
- The volume μ defined by $\mu(\Gamma) = \int_{\mathcal{X}} \mathbb{1}_{x \in \Gamma} dx$.
- \mathbb{Q} is taken uniform on \mathcal{A} .
- the γ_a , called test sets, defined through the scalar (or real valued vector) $a \in \mathbb{R}^m$:
For instance concentric disks of radius a or concentric squares of side a .

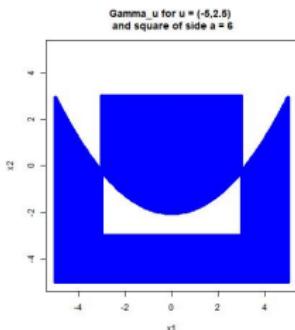
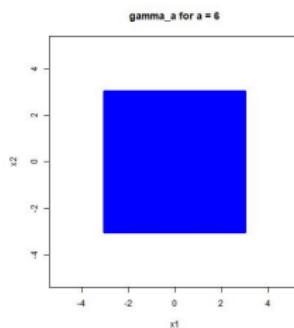
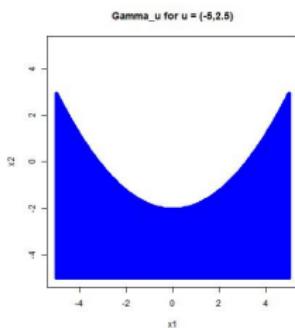
Idea of the universal index

$$S_{2, \text{ Univ}}^i (T_a, Q) := \frac{\int_A \text{Var}(\mathbb{E}[T_a(\Gamma_U) | U_i]) dQ(a)}{\int_A \text{Var}(T_a(\Gamma_U)) dQ(a)}$$



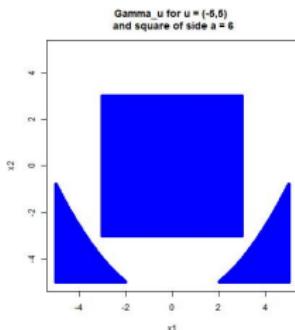
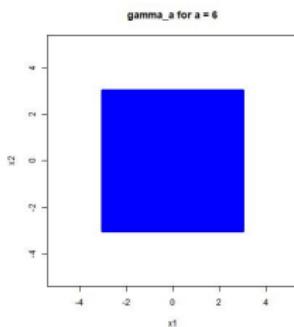
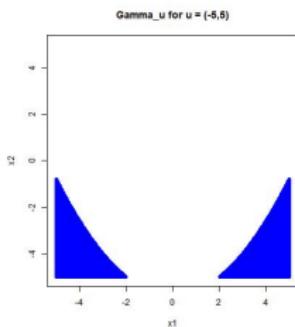
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Conclusion on universal indices on sets

Issues

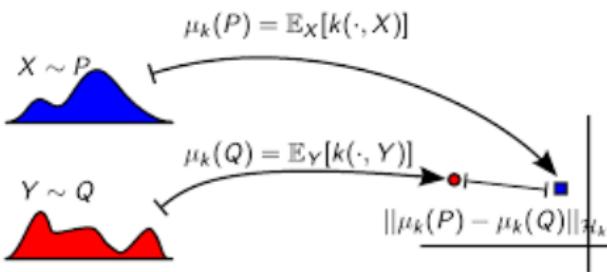
- Choice of T (the symmetric difference)
- Choice of \mathbb{Q}
- Choice of the test sets : we used squares and circles.

We could compare realizations to each other → Kernel-based sensitivity analysis.

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Distribution embedding into a RKHS



Distance between distributions as expectation of kernels, Gretton et al. 2006

$$\gamma_k^2(\mathbb{P}, \mathbb{Q}) := \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|_{\mathcal{H}_k}^2 = \mathbb{E}_{\substack{X, X' \sim \mathbb{P} \\ Y, Y' \sim \mathbb{Q}}} [k(X, X') + k(Y, Y') - 2k(X, Y)]$$

→ is a distance between distribution iff k is *characteristic* (i.e. injectivity of the embedding)

MMD-based index

MMD-based index

$$\begin{aligned} S_i &:= \mathbb{E}_{X_i} [\gamma_k^2(\mathbb{P}_Y, \mathbb{P}_{Y|X_i})] \\ &= \mathbb{E}_{X_i} \mathbb{E}_{\xi \sim \mathbb{P}_{Y|X_i}} \mathbb{E}_{\xi' \sim \mathbb{P}_{Y|X_i}} [k(\xi, \xi')] - \mathbb{E}_{Y \sim \mathbb{P}_Y} \mathbb{E}_{Y' \sim \mathbb{P}_Y} [k(Y, Y')]. \end{aligned}$$

→ has a ANOVA decomposition (daVeiga 2021) : can be used to rank the inputs by influence.

Estimation

- Pick & freeze estimation
- Rank-based estimation

HSIC-based index Gretton et al. 2006

HSIC-based index

$$\begin{aligned}
 HSIC_k(X_i, Y) &:= \gamma_k^2(\mathbb{P}_{X_i, Y}, \mathbb{P}_{X_i} \otimes \mathbb{P}_Y) \\
 &= \mathbb{E}_{X_i, X'_i, Y, Y'} k_{\mathcal{X}}(X_i, X'_i) k(Y, Y') \\
 &\quad + \mathbb{E}_{X_i, X'_i} k_{\mathcal{X}}(X_i, X'_i) \mathbb{E}_{Y, Y'} k(Y, Y') \\
 &\quad - 2 \mathbb{E}_{X_i, Y} [\mathbb{E}_{X'_i} k_{\mathcal{X}}(X_i, X'_i) \mathbb{E}_{Y'} k(Y, Y')]
 \end{aligned}$$

with (X'_i, Y') an independent copy of (X_i, Y) .

$HSIC_k(X_i, Y) = 0$ iif $X_i \perp Y$ (when k characteristic) → suited to identify the negligible inputs through independence testing.

Estimation

- Biased or unbiased classic estimators of $HSIC_k(X_i, Y)$
- p-values computable through asymptotic results or bootstrap methods

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SA on sets : a kernel between sets

Proposition (A kernel between sets, Balança et Herbin 2012)

The function $k_{set} : \mathcal{F}(\mathcal{X}) \times \mathcal{F}(\mathcal{X}) \rightarrow \mathbb{R}$ defined by

$$\forall \Gamma, \Gamma' \in \mathcal{F}(\mathcal{X}), \quad k_{set}(\Gamma, \Gamma') = \frac{\sigma^2}{2\lambda} e^{-\lambda \mu(\Gamma \Delta \Gamma')},$$

is positive definite for any positive scalars σ and λ .

Moore-Aronszajn theorem (Aronszajn (1950)) gives then the existence of a unique RKHS $\mathcal{H} \subset \mathcal{F}(\mathcal{X})^{\mathbb{R}}$ of reproducing kernel k_{set} .

Now we need to embed "random sets distributions" in \mathcal{H}

Random set distribution embedding

Definition (Capacity functional)

The capacity functional of a random closed set Γ denoted T_Γ is defined by :

$$T_\Gamma : \begin{array}{ccc} \mathcal{K}(\mathcal{X}) & \rightarrow & [0, 1] \\ K & \mapsto & \mathbb{P}(\Gamma \cap K \neq \emptyset). \end{array} \quad (7)$$

Definition (Mean embedding of a capacity functional)

The mean embedding of T_Γ is defined as

$$\mu_\Gamma = \mathbb{E}[k_{\text{set}}(\Gamma, \cdot)]. \quad (8)$$

Estimation with set-valued outputs

In the MMD- and HSIC-based indices expressions, we need to estimate quantities of the form $\mathbb{E}k_{set}(\Gamma_1, \Gamma_2)$.

With sets, $k_{set}(\Gamma_1, \Gamma_2)$ also require an estimation.

$$k_{set}(\Gamma_1, \Gamma_2) = e^{-\mu(\Gamma_1 \Delta \Gamma_2)} \quad (9)$$

$$= e^{-\mathbb{E}[\mathbb{1}_{\Gamma_1 \Delta \Gamma_2}(X)]} \text{ with } X \sim \mathcal{U}(\mathcal{X}) \quad (10)$$

$$\simeq e^{-\mu(\mathcal{X}) \frac{1}{N_x} \sum_{i=1}^{N_x} \mathbb{1}_{\Gamma_1 \Delta \Gamma_2}(x^i)} \quad (11)$$

$$= \widehat{k_{set}}(\Gamma_1, \Gamma_2) \quad (12)$$

Then we inject it in the indices estimators. For instance the normalized MMD-based index estimated through pick and freeze method is :

$$\widehat{S}_{A,p,f}^{MMD} = \frac{\sum_{i=1}^n \widehat{k_{set}}(\Gamma^{(i)}, \tilde{\Gamma}^{A,(i)}) - \widehat{k_{set}}(\Gamma^{(i)}, \Gamma'^{(i)})}{\sum_{i=1}^n \widehat{k_{set}}(\Gamma^{(i)}, \Gamma^{(i)}) - \widehat{k_{set}}(\Gamma^{(i)}, \Gamma'^{(i)})}. \quad (13)$$

Asymptotic behaviour of the indices on sets

Proposition (Quadratic error of a nested Monte Carlo estimator)

With the previous notations, using Rainforth et al. 2018, we have

$$\mathbb{E} \left(\frac{1}{n} \sum_{j=1}^n \widehat{k_{\text{set}}}(\Gamma_1^{(j)}, \Gamma_2^{(j)}) - \mathbb{E} k_{\text{set}}(\Gamma_1, \Gamma_2) \right)^2 = \mathcal{O}\left(\frac{1}{n} + \frac{1}{N_x^2}\right). \quad (14)$$

With this result, we can show that each quadratic error of our indices on sets has the same asymptotic behavior with rate $\mathcal{O}\left(\frac{1}{n} + \frac{1}{N_x^2}\right)$

Toy function, El-Amri et al. 2021

$$\forall x, u \in [-5, 5]^2 \times [-5, 5]^2 \quad g(x, u) = -x_1^2 + 5x_2 - u_1 + u_2^2 - 1. \quad (15)$$

$$\Gamma_U = \{x \in \mathcal{X}, g(x, U) \leq 0\}, \quad (16)$$

Toy function	Index	U_1	U_2
$g = -x_1^2 + 5x_2 - u_1 + u_2^2 - 1$	$\widehat{S_{i,p,f}^{MMD}}$	0.074	0.458
	$\widehat{S_{i,p,f}^{T,MMD}}$	0.545	0.935

Table – MMD-based indices

Toy function	Index	U_1	U_2
$g = -x_1^2 + 5x_2 - u_1 + u_2^2 - 1$	$\widehat{S_i^{HSIC_u}}$	0.072	0.284
	p-value	$4.70e-03$	$5.97e-10$

Table – HSIC-based indices

Test case : oscillator, Cousin 2021

$$g_1(\mathbf{x}, \mathbf{u}) = u_{r1} - \max_{t \in [0, T]} \mathcal{Y}'(x_1 + u_1, x_2 + u_2, u_p; t), \quad (17)$$

$$g_2(\mathbf{x}, \mathbf{u}) = u_{r2} - \max_{t \in [0, T]} \mathcal{Y}''(x_1 + u_1, x_2 + u_2, u_p; t), \quad (18)$$

with \mathcal{Y} the solution of the harmonic oscillator defined by :

$$(x_1 + u_1)\mathcal{Y}''(t) + u_p\mathcal{Y}'(t) + (x_2 + u_2)\mathcal{Y}(t) = \eta(t). \quad (19)$$

Uncertainty	Distribution	Uncertainty	Distribution
U_1	$\mathcal{U}[-0.3, 0.3]$	U_{r1}	$\mathcal{N}(1, 0.1^2)$
U_2	$\mathcal{U}[-1, 1]$	U_{r2}	$\mathcal{N}(2.5, 0.25^2)$
U_p	$\mathcal{U}[0.5, 1.5]$	U_{r3}	$\mathcal{N}(15, 3^2)$

Table – Definition of the uncertain inputs

Results on the oscillator case : MMD-based index

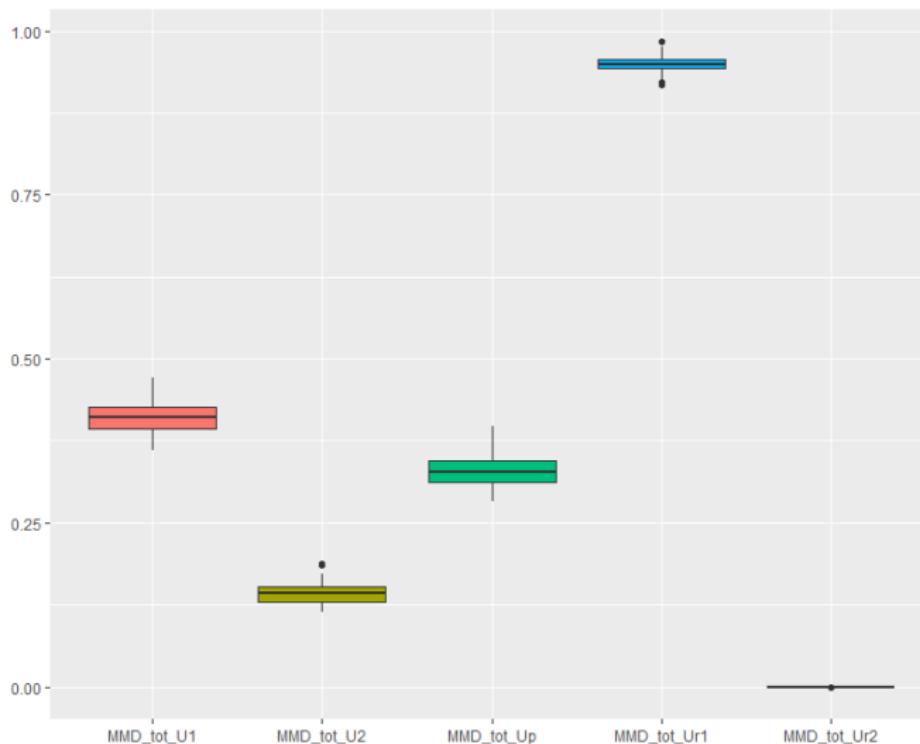


Figure – MMD-based total index for the constraint $g_1 \leq 0$

Results on the oscillator case : MMD-based index

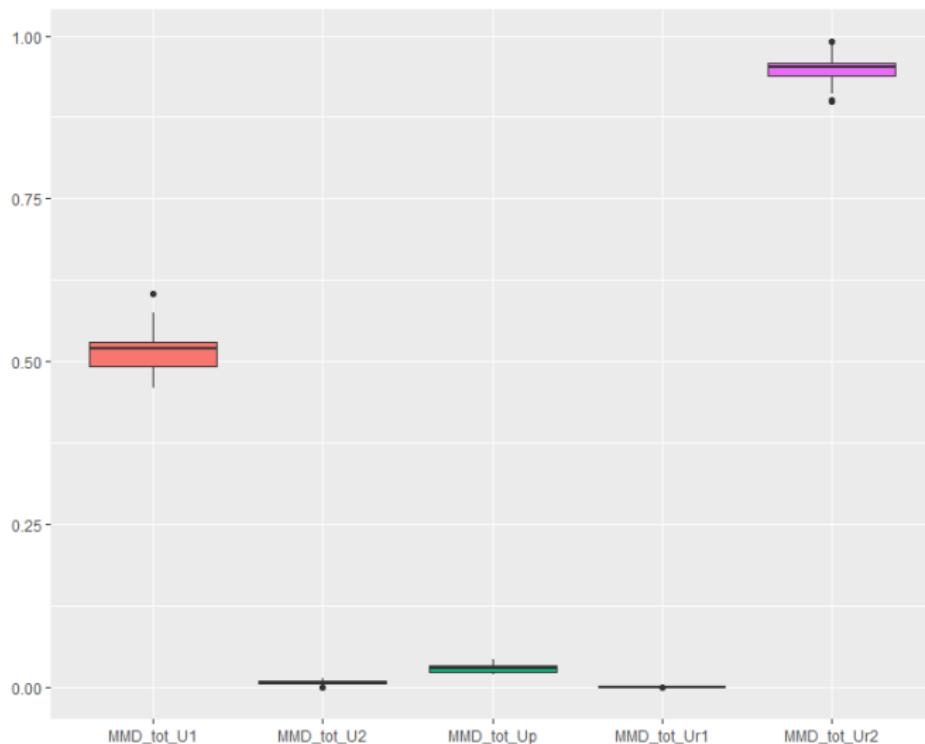


Figure – MMD-based total index for the constraint $g_2 \leq 0$

Results on the oscillator case : MMD-based index

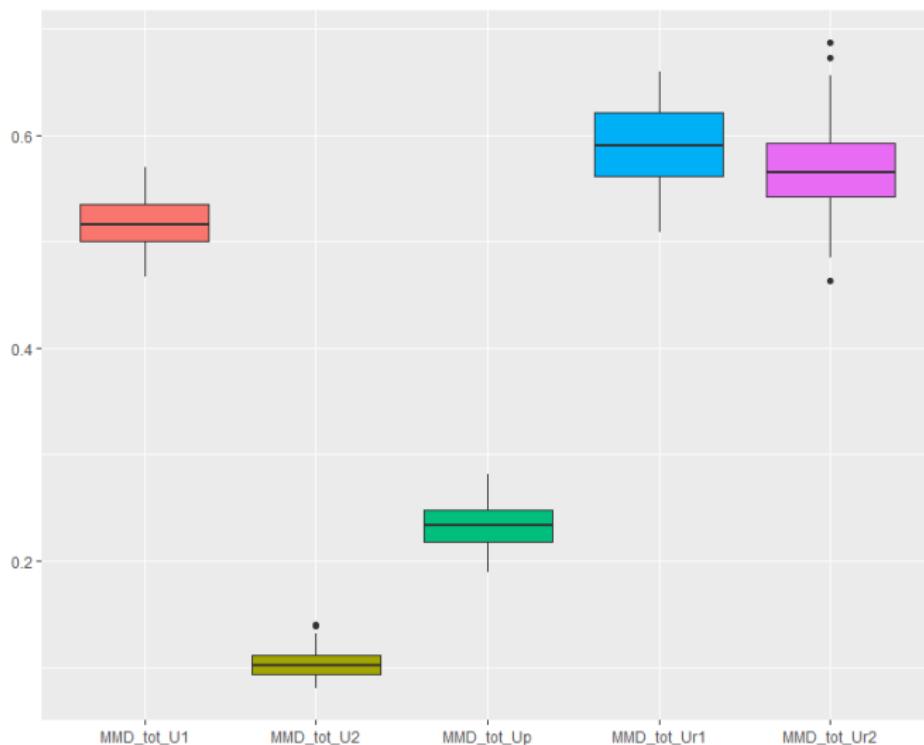


Figure – MMD-based total index for the constraint $g_1 \leq 0$ and $g_2 \leq 0$

Results on the oscillator case : MMD-based index

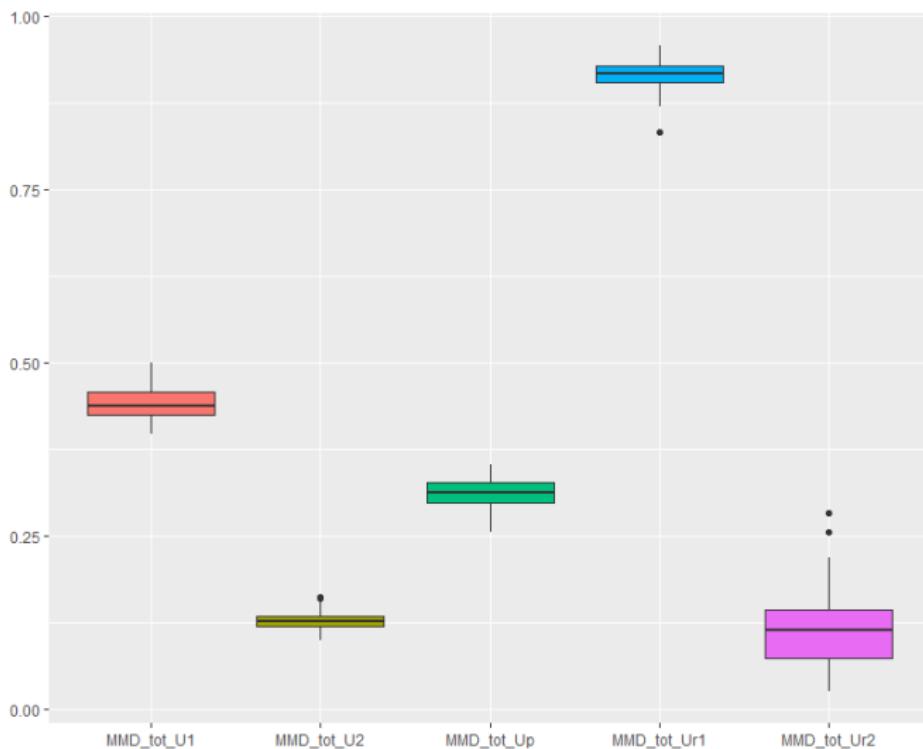


Figure – MMD-based total index for the couple ($g_1 \leq 0, g_2 \leq 0$)

Results on the oscillator case : HSIC-based index

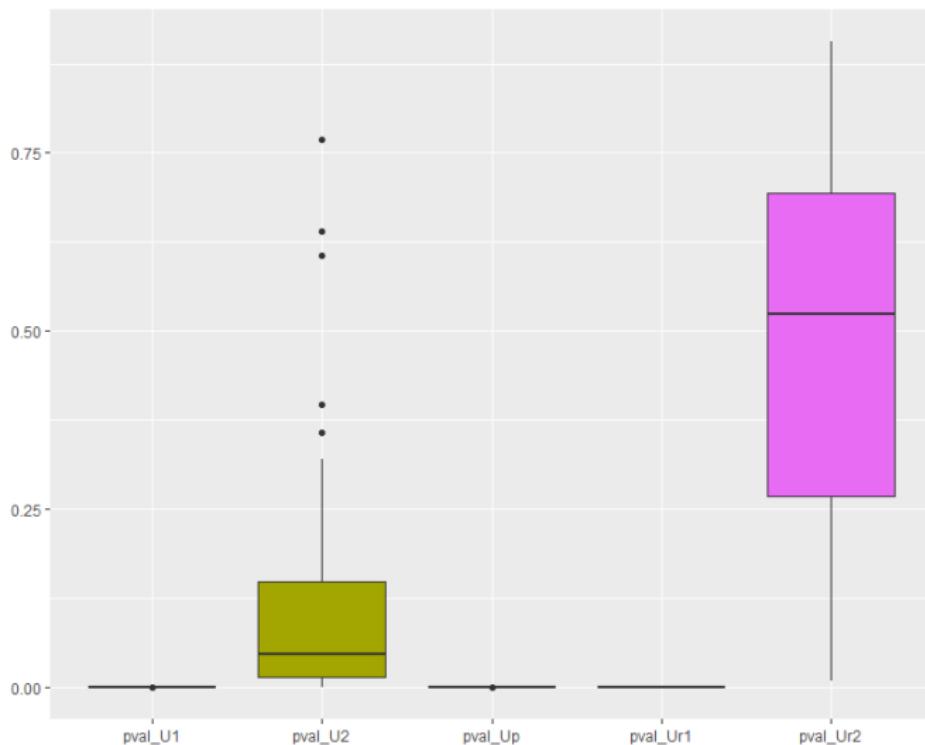


Figure – p-value of the HSIC-based index for the constraint $g_1 \leq 0$

Results on the oscillator case : HSIC-based index

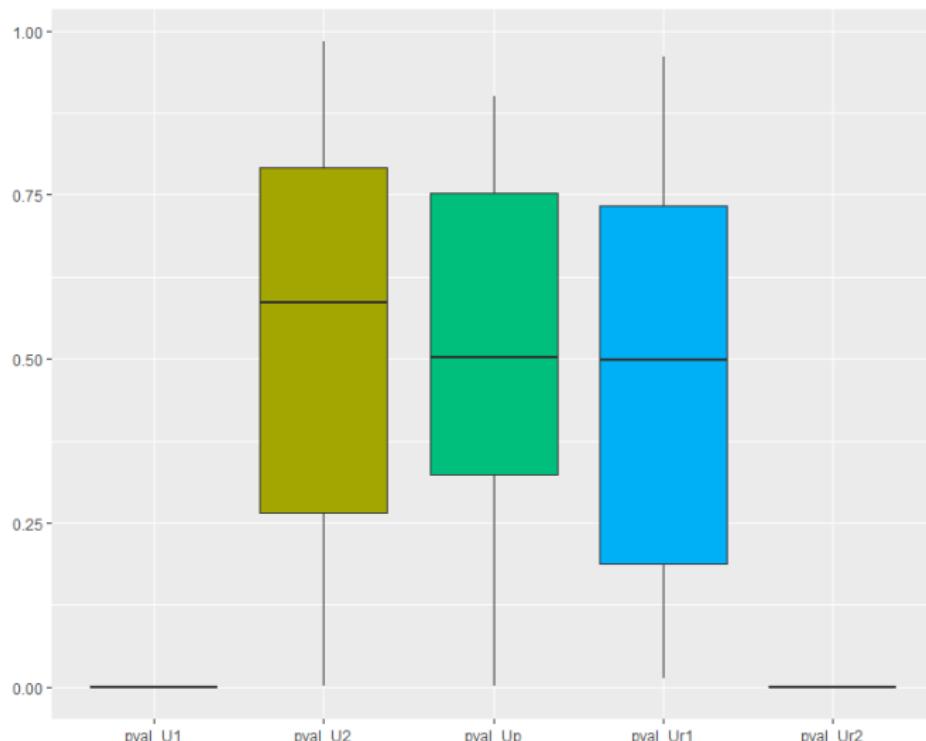


Figure – p-value of the HSIC-based index for the constraint $g_2 \leq 0$

Results on the oscillator case : HSIC-based index

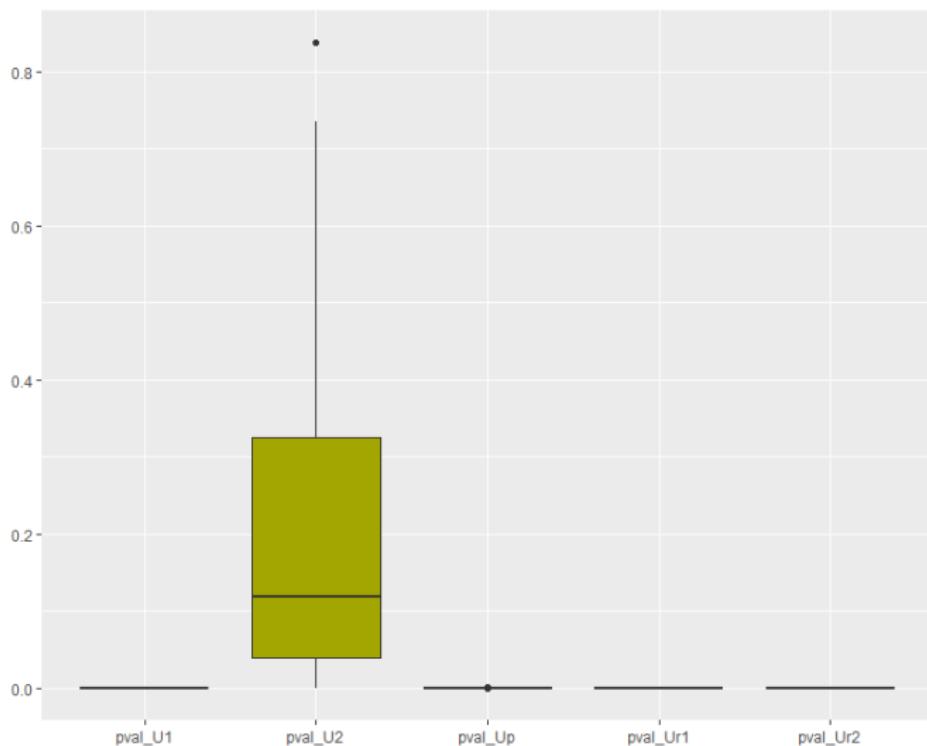


Figure – p-value of the HSIC-based index for the constraint $g_1 \leq 0$ and $g_2 \leq 0$

Results on the oscillator case : HSIC-based index

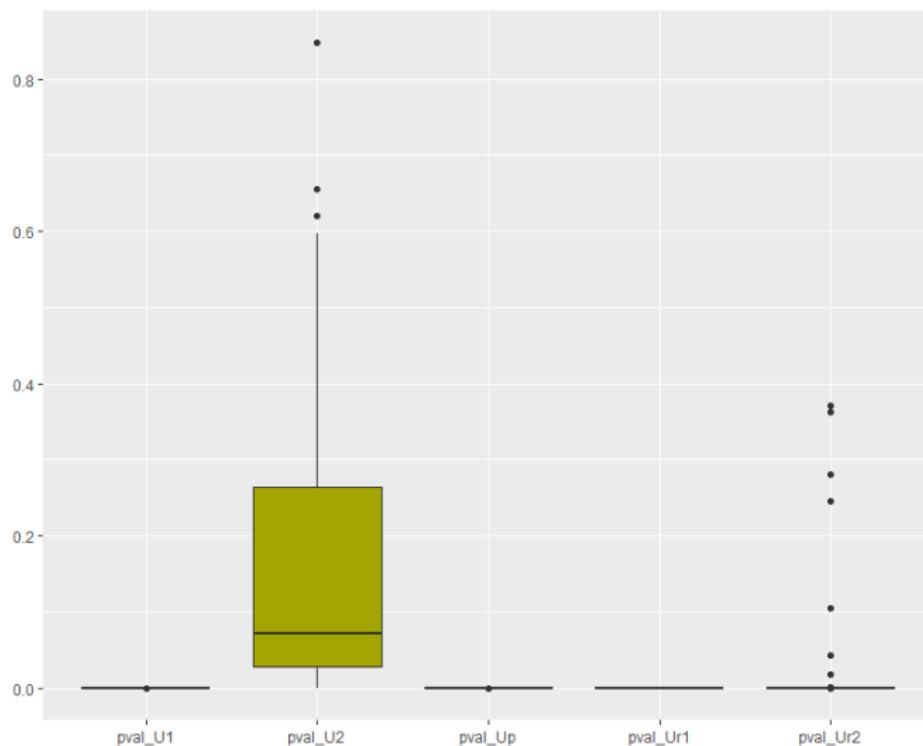


Figure – p-value of the HSIC-based index for the couple $(g_1 \leq 0, g_2 \leq 0)$

Conclusion

Kernel-based SA on set-valued outputs (paper "soon")

- A way to do SA on set-valued outputs
- On excursion sets : An answer to "How to do SA on the uncertain inputs in the context of robust optimization?"

Future work

- Test the three methods on a real test case (of Adan Reyes Reyes from IFPEN)
- Use it inside an optimization

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Excursion sets are random sets

For K compact, $h(U) = X = \{x \in \mathcal{X}, g(x, U) \leq 0\}$

$$\begin{aligned}\{K \cap X \neq \emptyset\} &=^c \{\omega, K \cap X(\omega) = \emptyset\} \\ &=^c \{\omega, \forall x \in K \ g(x, U(\omega)) > 0\} \\ &=^c \{\omega, \inf_{x \in K} g(x, U(\omega)) > 0\} \text{ as } K \text{ compact and } g \text{ continuous in } x \\ &=^c U^{-1}(\inf_{x \in K} g(x, \cdot)^{-1}(]0, +\infty[)) \in \mathcal{F}\end{aligned}$$

Indices estimation

$$S_{A,p.f}^{MMD} = \frac{\sum_{i=1}^n k_{set}(\Gamma^{(i)}, \tilde{\Gamma}^{A,(i)}) - k_{set}(\Gamma^{(i)}, \Gamma'^{(i)})}{\sum_{i=1}^n k_{set}(\Gamma^{(i)}, \Gamma^{(i)}) - k_{set}(\Gamma^{(i)}, \Gamma'^{(i)})}. \quad (20)$$

$$S_{I,rank}^{MMD} = \frac{\frac{1}{n} \sum_{i=1}^n k_{set}(\Gamma^{(i)}, \Gamma^{\sigma'_n(i)}) - \frac{1}{n^2} \sum_{i,j=1}^n k_{set}(\Gamma^{(i)}, \Gamma^{(j)})}{\frac{1}{n} \sum_{i=1}^n k_{set}(\Gamma^{(i)}, \Gamma^{(i)}) - \frac{1}{n^2} \sum_{i,j=1}^n k_{set}(\Gamma^{(i)}, \Gamma^{(j)})}. \quad (21)$$

$$\text{HSIC}_u(U_A, \Gamma) = \frac{1}{n(n-1)} \sum_{i,j=1, i \neq j}^n \left(k_{\mathcal{U}_A} \left(U_A^{(i)}, U_A^{(j)} \right) - 1 \right) k_{set} \left(\Gamma^{(i)}, \Gamma^{(j)} \right), \quad (22)$$

$$\text{HSIC}_b(U_A, \Gamma) = \frac{1}{n^2} \sum_{i,j=1}^n \left(k_{\mathcal{U}_A} \left(U_A^{(i)}, U_A^{(j)} \right) - 1 \right) k_{set} \left(\Gamma^{(i)}, \Gamma^{(j)} \right). \quad (23)$$